

PAC-Bayes-Chernoff bounds for unbounded losses

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PAC-Bayes theory provides high-probability generalization bounds for randomized learning algorithms.

- A randomized learning algorithms defines probability measure $\rho \in \mathcal{M}_1(\Theta)$ over the set of candidate models Θ .

Notation:

- Data, $D = \{\mathbf{x}_i\}_{i=1}^n$, is i.i.d. generated from an unknown distribution, ν , with support on $\mathcal{X}$
- We have a loss function $\ell : \Theta \times \mathcal{X} \rightarrow \mathbb{R}_+$
- *Population risk* of $\theta \in \Theta$ is defined as $L(\theta) := \mathbb{E}_\nu[\ell(\theta, \mathbf{X})]$
- *Empirical risk* of $\theta \in \Theta$ is defined as $\hat{L}(\theta, D) := \frac{1}{n} \sum_{i=1}^n \ell(\theta, \mathbf{x}_i)$

Standard PAC-Bayes bounds for bounded losses (McAllester, 2003):

$$\mathbb{E}_\rho[L(\theta)] \leq \mathbb{E}_\rho[\hat{L}(D, \theta)] + \sqrt{\frac{KL(\rho|\pi) + \log \frac{2\sqrt{n}}{\delta}}{2n}},$$

- The inequality holds simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$ with probability no less than $1 - \delta$ over the choice of $D \sim \nu^n$.
- We can minimize the bound with respect to ρ to obtain novel learning algorithms.

Unbounded losses are widely used in machine learning (e.g., cross-entropy, MSE).

- PAC-Bayes bounds for unbounded losses involve extra difficulties.
- Most existing PAC-Bayes bounds for unbounded losses are derived from next result:

PAC-Bayes (oracle) bound for unbounded losses

For any $\delta \in (0, 1)$ and any $\lambda > 0$, with probability at least $1 - \delta$ over draws of $D \sim \nu^n$,

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \frac{1}{\lambda} \left[\frac{\text{KL}(\rho|\pi) + \log \frac{f_{\pi, \nu}(\lambda)}{\delta}}{n} \right],$$

where $f_{\pi, \nu}(\lambda) := \mathbb{E}_{\pi} \mathbb{E}_{\nu^n} \left[e^{\lambda n (L(\boldsymbol{\theta}) - \hat{L}(D, \boldsymbol{\theta}))} \right]$.

[Alquier, P., Ridgway, J., & Chopin, N. (2016). On the properties of variational approximations of Gibbs posteriors. Journal of Machine Learning Research, 17(236), 1-41.]

Main difficulties:

- The exponential moment term $f_{\pi, \nu}(\lambda)$ has to be bounded using extra assumptions on the loss (e.g., sub-Gaussian assumption).
- The free parameter $\lambda > 0$ cannot be exactly optimized (discrete grid + union bounds).

Contributions

- A **novel PAC-Bayes oracle bound** for unbounded losses
 - Extends classic Cramér-Chernoff bounds to the PAC-Bayesian setup.
- Provides a **general framework to obtain empirical bounds** where:
 - **The free parameter λ is exactly optimized** without resorting to union-bound approaches.
 - The exponential moment term is averaged by the posterior, resulting in **more informative generalization bounds**.
 - Can be minimized to obtain **novel posteriors**.
- We illustrate the framework in several cases: generalized sub-Gaussian losses, L2 regularization, and input-gradient regularization.

Definition (Expected Cramér transform)

We define our *comparator* function for any $\rho \in \mathcal{M}_1(\Theta)$ as the *expected Cramér transform*:

$$\Lambda_{\rho}^{\star}(a) := \sup_{\lambda \in [0, b)} \{ \lambda a - \mathbb{E}_{\rho}[\Lambda_{\boldsymbol{\theta}}(\lambda)] \}, \quad a \in \mathbb{R}. \quad (1)$$

Where $\Lambda_{\boldsymbol{\theta}}(\lambda) := \log \mathbb{E}_{\nu} \left[e^{\lambda (L(\boldsymbol{\theta}) - \ell(\mathbf{x}, \boldsymbol{\theta}))} \right]$ is the Cumulant Generating Function (CGF) of the loss.

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Theorem (PAC-Bayes-Chernoff bound)

For any $\delta \in (0, 1)$, with probability at least $1 - \delta$ over draws of $D \sim \nu^n$,

$$\mathbb{E}_\rho[L(\theta)] \leq \mathbb{E}_\rho[\hat{L}(D, \theta)] + (\Lambda_\rho^*)^{-1} \left(\frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{n-1} \right),$$

simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$.

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simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$.

Equivalently:

$$\mathbb{E}_\rho[L(\theta)] \leq \mathbb{E}_\rho[\hat{L}(D, \theta)] + \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{\lambda(n-1)} + \frac{\mathbb{E}_\rho[\Lambda_\theta(\lambda)]}{\lambda}$$

simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$ and $\lambda \in (0, b)$.

Lemma

For any $\theta \in \Theta$ and $c \geq 0$, we have

$$\mathbb{P}_{D \sim \nu^n} \left(n\Lambda_{\theta}^* (L(\theta) - \hat{L}(D, \theta)) \geq c \right) \leq \mathbb{P}_{X \sim \exp(1)} (X \geq c).$$

Proof.

Careful rewriting of the Cramér-Chernoff bound after changes of variables. □

Using the fact that for any random variable Z with support $\Omega \subseteq \mathbb{R}_+$ its expectation can be written as

$$\mathbb{E}[Z] = \int_{\Omega} P(Z \geq z) dz, \tag{2}$$

the previous lemma will allow us to bound the exponential term $\mathbb{E}_{\nu^n} \left(e^{m\Lambda_{\theta}^* (L(\theta) - \hat{L}(D, \theta))} \right)$ in our main theorem.

Let $m < n$. First, by Jensen,

$$m\Lambda_{\rho}^{\star} \left(\mathbb{E}_{\rho} L(\boldsymbol{\theta}) - \mathbb{E}_{\rho} \hat{L}(D, \boldsymbol{\theta}) \right) \leq m \mathbb{E}_{\rho} \left[\Lambda_{\boldsymbol{\theta}}^{\star} (L(\boldsymbol{\theta}) - \hat{L}(D, \boldsymbol{\theta})) \right]$$

Apply Donsker-Varadhan's lemma on the right side to obtain

$$m\Lambda_{\rho}^{\star} \left(\mathbb{E}_{\rho} L(\boldsymbol{\theta}) - \mathbb{E}_{\rho} \hat{L}(D, \boldsymbol{\theta}) \right) \leq \text{KL}(\rho|\pi) + \log \mathbb{E}_{\pi} \left(e^{m\Lambda_{\boldsymbol{\theta}}^{\star} (L(\boldsymbol{\theta}) - \hat{L}(D, \boldsymbol{\theta}))} \right)$$

After Markov's inequality + Fubini, with probability at least $1 - \delta$,

$$m\Lambda_{\rho}^{\star} \left(\mathbb{E}_{\rho} L(\boldsymbol{\theta}) - \mathbb{E}_{\rho} \hat{L}(D, \boldsymbol{\theta}) \right) \leq \text{KL}(\rho|\pi) + \log \frac{1}{\delta} + \log \mathbb{E}_{\pi} \mathbb{E}_{\nu^n} \left(e^{m\Lambda_{\boldsymbol{\theta}}^{\star} (L(\boldsymbol{\theta}) - \hat{L}(D, \boldsymbol{\theta}))} \right).$$

Using the auxiliary lemma we obtain

$$\mathbb{E}_{\nu^n} \left(e^{m\Lambda_{\boldsymbol{\theta}}^{\star} (L(\boldsymbol{\theta}) - \hat{L}(D, \boldsymbol{\theta}))} \right) \leq \frac{n}{n - m}.$$

Taking $m = n - 1$ and applying $(\Lambda_{\rho}^{\star})^{-1}$ in both sides concludes the proof.

Let us unwrap our bound:

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \inf_{\lambda \in [0, b)} \left\{ \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{\lambda(n-1)} + \frac{\mathbb{E}_{\rho}[\Lambda_{\boldsymbol{\theta}}(\lambda)]}{\lambda} \right\}$$

simultaneously for every $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$.

Observations:

- Parameter-free bound without union-bounds at a $\log n$ cost.
- Generalization of $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$ depends on a three-way trade-off.
 - Minimize the empirical Gibbs loss $\mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})]$.
 - Minimize the KL term $\text{KL}(\rho|\pi)$.
 - **[Novel term]** Minimize the CGF term $\mathbb{E}_{\rho}[\Lambda_{\boldsymbol{\theta}}(\lambda)]$
- Directly related to regularization ([Masegosa&Ortega, 2023](#)).

Relation to previous bounds:

- If the loss is the 0-1 loss, we recover Langford-Seeger's bound ([Seeger, 2002](#)).

Bounded CGF assumptions is the standard approach to derived PAC-Bayes bounds for unbounded loss:

- If loss is σ -**sub-gaussian**,

$$\Lambda_{\boldsymbol{\theta}}(\lambda) \leq \frac{1}{2}\sigma^2\lambda^2 \quad \forall \boldsymbol{\theta} \in \Theta$$

- We can recover previous bounds ([Hellstrom & Durisi, 2021](#))

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \sqrt{2\sigma^2 \frac{KL(\rho|\pi) + \log \frac{n}{\delta}}{n-1}},$$

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simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$.

- Uniformly bound $\Lambda_{\boldsymbol{\theta}}(\lambda) \leq \psi(\lambda)$ for every $\boldsymbol{\theta} \in \Theta$, necessarily discards information about the statistical properties of individual models.
 - There are models with very different CGF ([Masegosa&Ortega, 2023](#)).

Definition (Model-dependent bounded CGF)

A loss function ℓ has model-dependent bounded CGF if for each $\boldsymbol{\theta} \in \Theta$, there is a convex and continuously differentiable function $\psi(\boldsymbol{\theta}, \lambda)$ such that $\psi(\boldsymbol{\theta}, 0) = \psi'(\boldsymbol{\theta}, 0) = 0$ and $\forall \lambda \geq 0$,

$$\Lambda_{\boldsymbol{\theta}}(\lambda) := \log \mathbb{E}_{\nu} \left[e^{\lambda (L(\boldsymbol{\theta}) - \ell(\mathbf{x}, \boldsymbol{\theta}))} \right] \leq \psi(\boldsymbol{\theta}, \lambda). \quad (3)$$

Theorem

Let ℓ be a loss function with model-dependent bounded CGF. Let $\pi \in \mathcal{M}_1(\Theta)$ be any prior independent of D . Then, for any $\delta \in (0, 1)$, with probability at least $1 - \delta$ over draws of $D \sim \nu^n$,

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{\lambda(n-1)} + \frac{\mathbb{E}_{\rho}[\psi(\boldsymbol{\theta}, \lambda)]}{\lambda}.$$

simultaneously for every $\rho \in \mathcal{M}_1(\Theta)$ and $\lambda \in (0, b)$.

If the loss function is σ^2 -sub-Gaussian, we have $\Lambda_{\boldsymbol{\theta}}(\lambda) \leq \frac{\lambda \sigma^2}{2}$ for every $\boldsymbol{\theta} \in \Theta$ (Hellström & Durisi, 2021):

$$\mathbb{E}_{\rho} L(\boldsymbol{\theta}) \leq \mathbb{E}_{\rho} \hat{L}(D, \boldsymbol{\theta}) + \sqrt{2 \sigma^2 \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{n-1}}. \quad (4)$$

Our theorem allows the following, more general assumption: $\Lambda_{\boldsymbol{\theta}}(\lambda) \leq \frac{\lambda \sigma(\boldsymbol{\theta})^2}{2}$ for each $\boldsymbol{\theta} \in \Theta$.

$$\mathbb{E}_{\rho} L(\boldsymbol{\theta}) \leq \mathbb{E}_{\rho} \hat{L}(D, \boldsymbol{\theta}) + \sqrt{2 \mathbb{E}_{\rho}[\sigma(\boldsymbol{\theta})^2] \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{n-1}}, \quad (5)$$

Remark:

The proxy variance σ^2 in equation (4) is a worst-case constant, hence (5) is more general and potentially tighter. It also shows that generalization depends on finding models with smaller proxy variance.

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However, **we are not limited to using tail assumptions!**

L2 regularization minimizes an objective function of the form

$$\hat{L}(D, \boldsymbol{\theta}) + k \|\boldsymbol{\theta}\|_2^2,$$

where $k > 0$ is a trade-off parameter.

If the loss $\ell(\mathbf{x}, \boldsymbol{\theta})$ is M -Lipschitz with respect to $\boldsymbol{\theta}$, as shown in (Masegosa&Ortega, 2023),

$$\Lambda_{\boldsymbol{\theta}}(\lambda) \leq 2M\lambda^2 \|\boldsymbol{\theta}\|_2^2. \quad (6)$$

Theorem

If ℓ satisfies the inequality above, then for any $\delta_1 \in (0, 1)$, with probability at least $1 - \delta_1$ over draws $D \sim \nu^n$,

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \sqrt{2M \mathbb{E}_{\rho}[\|\boldsymbol{\theta}\|_2^2] \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta_1}}{n-1}} \quad (7)$$

simultaneously for every $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$.

Input-gradient regularization ([Varga et al., 2017](#)) minimizes an objective function of the form

$$\hat{L}(D, \boldsymbol{\theta}) + k \frac{1}{n} \sum_{i=1}^n \|\nabla_{\mathbf{x}} \ell(\mathbf{x}_i, \boldsymbol{\theta})\|_2^2,$$

where $k > 0$ is a trade-off parameter. This approach is often used to make models more robust against disturbances in input data and adversarial attacks ([Ross & Doshi-Velez, 2018](#)). With the proper assumptions, our bounds provide a **PAC-Bayesian interpretation**.

The connection is provided by the assumption that the underlying distribution satisfies a log-Sobolev inequality ([Chafai, 2004](#)):

$$\Lambda_{\boldsymbol{\theta}}(\lambda) \leq \frac{M}{2} \lambda^2 \mathbb{E}_{\nu} \|\nabla_{\mathbf{x}} \ell(\mathbf{x}, \boldsymbol{\theta})\|_2^2 \quad (8)$$

for every $\lambda > 0$ and some $M > 0$.

Theorem (Oracle PAC-Bayes bound for input-gradients)

If ν satisfies the inequality above, then for any $\delta_1 \in (0, 1)$, with probability at least $1 - \delta_1$ over draws $D \sim \nu^n$,

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \sqrt{2M \mathbb{E}_{\rho} \left[\mathbb{E}_{\nu} \|\nabla_{\mathbf{x}} \ell(\mathbf{x}, \boldsymbol{\theta})\|_2^2 \right] \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta_1}}{n-1}} \quad (9)$$

simultaneously for every $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$.

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simultaneously for every $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$.

We can obtain an empirical bound if we assume that the gradient norms have sub-Gaussian tails.

Theorem (Empirical PAC-Bayes bound for input-gradients)

With the same conditions as above, assume $\|\nabla_{\mathbf{x}} \ell(\mathbf{x}, \boldsymbol{\theta})\|_2^2$ is σ^2 -sub-Gaussian for every $\boldsymbol{\theta} \in \boldsymbol{\Theta}$. Then for any $\delta \in (0, 1)$, with probability at least $1 - \delta$ over draws of $D \sim \nu^n$,

$$\begin{aligned} \mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] &+ \sqrt{2M \left(\frac{1}{n} \sum_{i=1}^n \mathbb{E}_{\rho} [\|\nabla_{\mathbf{x}} \ell(\mathbf{x}_i, \boldsymbol{\theta})\|_2^2] + \frac{\sigma^2}{2} \right) \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{n-1}} \\ &+ \sqrt{2M \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{n-1}}. \end{aligned}$$

simultaneously for every $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$.

If we fix $\lambda > 0$ and repeat the same procedure —log-Sobolev + sub-Gaussianity of gradients— starting with the parametric bound

$$\mathbb{E}_{\rho}[L(\boldsymbol{\theta})] \leq \mathbb{E}_{\rho}[\hat{L}(D, \boldsymbol{\theta})] + \frac{\text{KL}(\rho|\pi) + \log \frac{n}{\delta}}{\lambda(n-1)} + \frac{\mathbb{E}_{\rho}[\Lambda_{\boldsymbol{\theta}}(\lambda)]}{\lambda},$$

the subsequent bound can be minimized w.r.t. $\rho \in \mathcal{M}_1(\boldsymbol{\Theta})$, resulting in the optimal posterior:

$$\rho^*(\boldsymbol{\theta}) \propto \pi(\boldsymbol{\theta}) \exp \left\{ - (n-1) \left(\hat{L}(D, \boldsymbol{\theta}) + k \frac{1}{n} \sum_{i=1}^n \|\nabla_{\mathbf{x}} \ell(\mathbf{x}_i, \boldsymbol{\theta})\|_2^2 \right) \right\},$$

where $k = \lambda \frac{M}{2}$. The optimal posterior concentrates its mass of probability in models minimizing the term $\hat{L}(D, \boldsymbol{\theta}) + \frac{\lambda M}{2} \frac{1}{n} \sum_{i=1}^n \|\nabla_{\mathbf{x}} \ell(\mathbf{x}_i, \boldsymbol{\theta})\|_2^2$, which is exactly the minimization objective of input-gradient regularization with trade-off parameter $\frac{\lambda M}{2}$.

- **Novel PAC-Bayes Oracle Bound:** Introduced a novel PAC-Bayes oracle bound leveraging the Cramer transform.
- **Optimization of Free Parameter** ($\lambda > 0$): Facilitates exact optimization of the free parameter $\lambda > 0$, solving a longstanding issue in PAC-Bayesian methods and allowing for tighter empirical bounds.
- **Introduction of Model-Dependent Assumptions:** Enables flexible, richer model-dependent assumptions for bounding the Cumulative Generating Function (CGF).
 - **Applicability Across Diverse Assumptions:** Demonstrates the bound's utility through assumptions such as generalized sub-Gaussian losses, parameter norm bounds, and log-Sobolev inequalities, including a novel empirical PAC-Bayes bound.

Thanks for your attention!! :)